

## Psychometric Properties of the Temperament and Character Inventory-Revised (TCI-R) in a Belgian Sample

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**Abstract:** The Temperament and Character Inventory (TCI; Cloninger, Przybeck, Svrakic, & Wetzell, 1994) is a self-questionnaire developed to assess the 7 dimensions of personality described by Cloninger et al. (1994) with a total of 29 subscales. In 1999, a revised version was proposed by Cloninger (TCI-R). In this study, we present psychometric properties of the TCI-R from 958 French-speaking participants of Belgium. Women exhibited higher scores for harm avoidance, reward dependence, and cooperativeness dimensions. The proposed factorial structure of 4 temperament dimensions and 3 character dimensions was confirmed. The TCI-R inventory had good test-retest reliabilities as well as good alpha coefficients. The addition of 3 new subscales to the original scale for Persistence has produced a very reliable dimension in the TCI-R.

Over the past decade, Cloninger and his colleagues (Cloninger, 1986, 1987; Cloninger & Svrakic, 1997; Cloninger, Svrakic, & Przybeck, 1993; Svrakic, Whitehead, Przybeck, & Cloninger, 1993) developed a biosocial model of personality based on four temperaments (novelty seeking [NS], harm avoidance [HA], reward dependence [RD], and persistence [PS]) and three characters (self-directedness [SD], cooperativeness [C], and self-transcendence [ST]). According to Cloninger (Cloninger et al., 1993), *temperaments* refer to automatic emotional responses to experiences that are moderately heritable and stable throughout life. In contrast, *characters* refer to self-concepts and individual differences in goals and values. Character is moderately influenced by insight and learning. It is unique to each individual, and it matures in progressive steps. The model of personality proposed by Cloninger (Cloninger et al., 1993) is distinctive in that it is rationally based on robust findings about neurobiological and experiential influences on personality structure and development. Theoretically and empirically, the relations among temperament and character dimensions are strongly nonlinear, reflecting the hierarchical nature of the supervisory mental processes (character) that regulate emotional conflicts (temperament; Cloninger, Svrakic, & Svrakic, 1997).

Briefly, the four temperaments and three characters can be described as follows. NS is the tendency to respond actively to novel stimuli leading to pursuit of rewards and escape from punishment. HA is the tendency to inhibit responses to signals of aversive stimuli that lead to avoidance of punishment and nonreward. RD is the tendency for a positive response to conditioned signals of reward that maintain behavior. PS is perseverance despite frustration and fatigue based on resistance to extinction of intermittently reinforced behavior. SD is the ability of an individual to control, regulate, and adapt his or her behavior to fit the situation in accord with individually chosen goals and values. C accounts for individual differences in identification with and acceptance of other people. ST is associated with spirituality, and it refers generally to identification with everything conceived as essential and consequential parts of a unified whole. When associated with higher scores on SD and C, higher scores on ST refer to creative character.

Many psychometric and neurobiological investigations have described results that contribute to the validation of the model (Cloninger, Przybeck, & Svrakic, 1991; Hansenne & Ansseau, 1999; Hansenne et al., 2002; Hennig, Toll, Schonlau, Rohrmann, & Netter, 2000; Svrakic, Przybeck, & Cloninger, 1991). According to Cloninger's (1986, 1987) model, monoamine neurotransmitter systems are considered as the bases of fundamental personality traits (Cloninger, 1986, 1987). Each of the three fundamental temperaments (i.e., NS, HA, and RD) has been associated with a specific central neurotransmitter: NS with low basal dopaminergic activity, HA with high serotonergic activity, and RD with low basal noradrenergic activity. The association between NS and dopaminergic function has been validated in several studies (e.g., Gerra et al., 2000; Hansenne et al., 2002; Wiesbeck, Mauerer, Thome, Jacob, & Boening, 1995). However, some studies have not reported a relationship between NS and dopaminergic activity (e.g., Hansenne & Ansseau, 1998; Limson et al., 1991). In the same vein, some studies have demonstrated a relationship between norepinephrine and RD (Curtin et al., 1997; Garvey, Noyes, Cook, & Blum, 1996), whereas others have failed to report such associations (Hansenne & Ansseau,

1998; Limson et al., 1991). Concerning the relationship between serotonergic function and HA, some studies have provided data that validate the theoretical association (Hansenne & Ansseau, 1999; Hansenne et al., 1997).

The Temperament and Character Inventory (TCI; Cloninger, Przybeck, Svrakic, & Wetzel, 1994) is a 226-item self-questionnaire developed by Cloninger et al. to assess the seven dimensions of personality. The variances explained by the rotated factors after promax rotation ranged from 4.4% to 2.0%, which is 17.7% to 8.0% of the total variance in the 25 factors (Cloninger et al., 1993). Cloninger et al. (1993) reported significant relationships among the seven dimensions: C was positively correlated with both RD (.54) and SD (.57), whereas SD was negatively correlated with HA (-.47). Weaker correlations were found between PS and HA (-.27), between C and HA (-.28), between SD and PS (.28), and between ST and RD (.28). Hansenne, Le Bon, Gauthier, and Ansseau (2001) reported similar correlations within a Belgian sample. Many psychometric studies have confirmed the structure of the Cloninger's (Cloninger et al., 1993) model. Pélissolo and Lépine (2000), as well as Hansenne et al. (2001), provided confirmatory findings from the French version of the TCI. In addition, several studies have validated the structure of the questionnaire (Brändström et al., 1998; Gutierrez et al., 2001; Tomina et al., 2000). However, Gana and Trouillet (2003) presented an unsuccessful attempt to validate the structure of the TCI.

The TCI provides quantitative measures of personality that are clinically useful in psychiatry and psychology (Cloninger & Svrakic, 1997). First, the TCI provides personality measures that quantify individual differences in vulnerability to many Axis I disorders (e.g., major depressive disorder, anxiety disorders, eating disorders, substance dependencies, and also psychoses). Indeed, several studies have demonstrated that depressed patients, even in remission states, have higher HA scores and that severity of depression correlates with HA (Corruble, Duret, Pélissolo, Falissard, & Guelfi, 2002; Grucza, Przybeck, Spitznagel, & Cloninger, 2003; Hansenne et al., 1999; Marijnissen, Tuinier, Sijben, & Verhoeven, 2002). Moreover, patients who fail to respond to antidepressant treatments generally have higher HA scores before treatment than the others. Overall, these results support various relationships between depression and HA: an influence of the depressive state on a trait measure, a vulnerability model (HA would be a susceptibility factor for depression), and a scar model with elevated HA scores even after remission of acute depressive symptoms. Other temperament dimensions, NS and RD, are not consistently associated with depressive status, but the character dimension of SD is often low when compared to participants without depressive disorders. Additional studies have found a predictive value of HA scores for suicidal attempts (Bulik, Sullivan, & Joyce, 1999; Engstrom, Brändström, Sigvardson, Cloninger, & Nylander, 2004). Furthermore, in a sample of schizophrenia patients (Eklund, Hansson, & Bengtsson-Tops, 2004), SD was highly correlated with psychological health. Moreover, Van Ammers, Sellman, and Mulder (1997) found a significant correlation between the NS dimension and use of alcohol, cannabis, caffeine, and nicotine in schizophrenia patients, suggesting that NS behaviors contribute to substance use in schizophrenia.

Second, the TCI is a reliable instrument to assess personality disorders: lower SD and C scores have been found consistently in individuals with personality disorders (Bayon, Hill, Svrakic, Przybeck, & Cloninger, 1996; Svrakic et al., 1993). Personality disorders have been shown to be characterized by low SD and low C regardless of the cluster or category of personality disorder (Svrakic et al., 1993). Therefore, these two dimensions have been proposed to be the core features of personality disorders. Besides this general finding, patients in clusters A (aloof), B (impulsive), and C (fearful) were differentiated respectively by low RD, high NS, and high HA. More specifically, high NS, low RD and low HA characterize antisocial personality disorder, whereas low NS, high RD, and high HA characterize avoidant personality disorder (Svrakic et al., 1993). In addition, SD was significantly correlated with personality disorders assessed by a Dutch questionnaire (de la Rie, Duijsens, & Cloninger, 1998). Moreover, significant negative correlations occurred between C and most personality disorders.

The TCI has been translated in many foreign languages used on all continents including Swedish, Dutch, Japanese, Spanish, and French (Brändström et al., 1998; de la Rie et al., 1998; Pélissolo & Lépine, 1997, 2000; Tanaka, Sakamoto, Kijima, & Kitamura, 1998). The French version of the TCI is well validated (Pélissolo & Lépine, 2000), and it is used in the French-speaking part of Belgium (Hansenne, 1999; Hansenne et al., 1999). Normative data from a Belgian sample were published (Hansenne et al., 2001). Interestingly, Belgian normative data are similar to French ones (Pélissolo & Lépine, 2000) but different from the American norms (Cloninger et al., 1993). Compared to American norms, both Belgian and French individuals exhibited lower scores on NS, PS, and ST dimensions, and conversely, Belgian and French individuals had higher scores on HA.

In recent times, Cloninger (1999) developed a revised version of the TCI (TCI-R) introducing two major modifications. First, the original TCI was a true-false questionnaire, but in the TCI-R, the participants must respond on a 5-point Likert scale ranging from 1 (*definitively false*) to 5 (*definitively true*) to enhance the precision of measurement for subscales. Second, in the original TCI, PS was measured by only one short scale,

whereas in the TCI-R, this dimension has 35 items and four subscales (eagerness of effort vs. laziness; work hardened vs. spoiled, ambitious vs. underachieving, perfectionist vs. pragmatist) to improve its description and measurement. In addition, the TCI-R adds a new subscale for RD (RD2; open to warm communication vs. aloofness).

Until recently, few data have been available about the TCI-R. Brändström, Richter, and Nylander (2003) demonstrated structural equivalence of the TCI and the TCI-R from a cross-cultural perspective among healthy participants from Sweden and Germany. Moreover, the reliability coefficients for the revised version were higher than those found for the TCI within both samples. Pélissolo et al. (in press) found a good factorial structure for the TCI-R and similar to those shown with the TCI in general and clinical populations. Robust factors were obtained for RD and PS in the TCI-R, even more clearly than in the TCI. All dimensions obtained higher alpha Cronbach coefficients with the TCI-R than with the TCI. Last, Gutierrez-Zotes et al. (2004) provided normative data of the Spanish version of the TCI-R and found higher reliability coefficients for all the dimensions. The aim of this study is to provide further psychometric validity of the TCI-R for the French TCI-R in a Belgian sample.

## MATERIALS AND METHOD

The study was conducted on a sample of 958 adults from the French part of Belgium. The sample was made up of 546 women with a mean age of 33.6 years (range from 17 to 70 years,  $SD = 15.1$ ) and 412 men with a mean age of 36.2 years (range from 17 to 70 years,  $SD = 14.9$ ). Among the participants, 352 were undergraduate students (284 women) who were enrolled in psychological courses. The other participants consisted of student's relatives. All participants gave informed consent to participate in the study.

A 2-week retest was proposed for a subsample of 95 participants, but only 77 filled out the TCI-R for a second time. This subsample comprised 68 women (mean age of 19.8,  $SD = 1.8$ ) and 9 men (mean age of 20.8,  $SD = 1.9$ ) who were all undergraduate students. Moreover, a 6-month retest was obtained in a different subsample of 40 participants (30 women with a mean age of 22.7,  $SD = 8.5$  and 10 men with a mean age of 31.8,  $SD = 15.7$ ). This subsample included undergraduate students (28) and relatives.

The participants filled out the French version of the 240 items of the TCI-R. Undergraduate students answered the questionnaire during a course, and the other participants responded at home. The translation was performed by Pélissolo, Notides, Musa, Téhérani, and Lépine (2000), who previously translated the TCI in French (Pélissolo & Lépine, 2000).

After the completion, one author (M. Hansenne) checked if the participants filled all the items and if they responded appropriately to the five validity items. A total of 26 participants were not included in the study according to incomplete responses (more than 5 lacking responses) and/or poor validity items (at least one poor validity item).

First, scale and subscale means and their distributions were examined, and Pearson's correlations were performed between scales as well as between the scales and age. Second, gender differences on TCI-R scales and subscales were examined with multivariate analysis of covariance, with TCI-R scores as dependent variables, gender as an independent variable, and age as cofactor. Cohen's  $d$  effect size measures were also computed after controlling for age. Third, test-retest reliabilities were assessed using the two-way random effects model intraclass correlation coefficient (ICC; 2,1; McGraw & Wong, 1996). Additionally, a one-way repeated measure analysis of variance (ANOVA) with the factor of time as well as Cohen's  $d$  effect size measures were calculated. Fourth, the internal consistency was estimated with the Cronbach's  $\alpha$  coefficient. The coefficients for the dimensions were derived from all the individual items (i.e., all NS items and not NS1, NS2, NS3, NS4 as items). Fifth, the structure of the TCI-R was evaluated by a principal component analysis rotating the factors by promax. Temperaments and characters were analyzed separately. The relations among the temperament dimensions are nearly linear and the relations among the character dimensions are also nearly linear. However, the relationships among the temperament and character dimensions are strongly nonlinear, which would create problems in a factor analysis examining all the dimensions simultaneously. Nonlinearity is shown by the fact that there are no one-to-one (linear) relationships among configurations of temperament with configurations of character: There is extensive equifinality (i.e., different temperaments are associated with the same character configuration) and multifinality (i.e., the same temperament configuration has different character outcomes; see Cloninger, 2000, 2004; Cloninger et al., 1997). For example, factor analysis would force HA and SD into a single factor because they are moderately correlated, especially in immature people, but they are in fact dissociable biologically and developmentally as shown by distinguishing between personality disorders and anxiety disorders (there are nonanxious personality disorders and anxious mature people).

The criterion for identifying the number of factors to extract was based on parallel analysis (Glorfeld, 1995; O'Connor, 2000). Indeed, the factor analysis literature has clearly indicated that researchers should not use the default decision rule of eigenvalues greater than 1 or examination of scree plots of eigenvalues. In parallel analysis, the rational is to select the number of components that account for more variance than the components derived from random data.

## RESULTS

The mean scores for TCI-R scales and their subscales are presented in Table 1. Multivariate analysis showed that women exhibited significantly higher scores for HA, RD, and C dimensions than men. For the C dimension, all the subscale scores were higher in women, whereas for the HA and RD dimensions, one subscale did not differ between the two groups (respectively, HA4 and RD4). Moreover, NS4 and PS3 subscale scores were lower in women, whereas SD4 was higher. Alpha coefficients ranged from 0.45 (RD4) to 0.90 (HA and PS), with a mean value of 0.70. Globally, the internal consistency of the scales was good, whereas the internal consistency of the subscales was not so good, with many of them below 0.70. Concerning the dimensions, NS had the lowest alpha value (0.78).

**TABLE 1:** TCI-R Scale and Subscale Means, Standard Deviations, Cronbach alpha, and Cohen's *d* Among the Sample

TCI-R Dimensions	Total		$\alpha$	Women		Men		F (df = 1952)	P	<i>d</i> <sup>a</sup>
	M	SD		M	SD	M	SD			
NS1 (exploratory excitability)	31.2	5.4	.63	31.4	5.3	30.8	5.5	0.71	.39	0.055
NS2 (impulsiveness)	24.0	5.3	.58	24.2	5.2	23.7	5.4	2.53	.11	0.103
NS3 (extravagance)	25.8	6.3	.78	26.1	6.6	25.6	6.0	1.43	.23	0.078
NS4 (disorderliness)	19.1	4.4	.57	18.5	4.1	20.0	4.6	40.30	<.001	0.412
Novelty seeking (NS)	100.2	14.2	.78	100.2	14.4	100.0	13.9	0.19	.67	0.028
HA1 (anticipatory worry)	30.6	6.2	.78	32.3	6.7	28.4	6.3	78.84	<.001	0.576
HA2 (fear of uncertain)	22.4	5.3	.74	23.7	5.2	20.7	5.1	87.74	<.001	0.607
HA3 (shyness with strangers)	20.6	6.1	.85	21.4	6.2	19.6	5.9	16.17	<.001	0.261
HA4 (fatigability)	20.4	5.3	.73	20.6	5.3	20.2	5.2	0.93	.34	0.063
Harm avoidance (HA)	94.0	18.2	.90	97.9	18.1	88.9	17.0	58.42	<.001	0.496
RD1 (sentimentality)	28.9	4.2	.57	29.9	4.1	27.6	4.0	74.76	<.001	0.561
RD2 (open to warm communication)	34.4	5.9	.72	35.5	6.0	32.9	5.8	43.27	<.001	0.426
RD3 (attachment)	19.8	4.9	.59	20.4	5.2	18.9	4.4	19.20	<.001	0.284
RD4 (dependence)	18.7	3.8	.45	18.8	3.8	18.4	3.7	1.83	.18	0.088
Reward dependence (RD)	101.7	13.4	.81	104.6	13.8	97.8	12.3	57.42	<.001	0.491
PS1 (eagerness of effort)	30.2	5.6	.76	30.4	5.3	29.9	6.0	5.46	.02	0.151
PS2 (work hardened)	27.9	5.0	.76	27.6	4.7	28.3	5.2	3.09	.07	0.114
PS3 (ambitious)	32.9	5.8	.73	32.3	5.8	33.7	5.7	13.50	<.001	0.238
PS4 (perfectionist)	26.7	5.3	.69	26.4	5.4	27.1	5.2	2.76	.09	0.108
Persistence (PS)	117.8	18.4	.90	116.8	18.2	119.1	18.7	2.07	.15	0.093
SD1 (responsibility)	29.1	5.1	.73	29.2	4.8	28.9	5.5	1.77	.18	0.086
SD2 (purposefulness)	21.8	4.0	.67	21.9	4.0	21.8	3.9	0.14	.70	0.024
SD3 (resourcefulness)	17.8	3.3	.60	17.6	3.2	18.1	3.4	4.17	.04	0.132
SD4 (self-acceptance)	33.2	7.3	.73	33.7	7.3	32.5	7.3	11.22	<.001	0.217
SD5 (congruent second nature)	38.1	5.8	.48	37.8	5.8	38.5	5.7	1.44	.23	0.078
Self-directedness (SD)	140.1	17.4	.81	140.3	17.4	139.9	17.4	1.20	.27	0.071
C1 (social acceptance)	30.5	4.5	.59	31.1	4.6	29.7	4.4	25.07	<.001	0.325
C2 (empathy)	18.2	3.1	.58	18.8	2.9	17.4	3.1	50.07	<.001	0.459
C3 (helpfulness)	30.9	4.0	.57	31.5	4.0	30.0	3.9	35.18	<.001	0.385
C4 (compassion)	25.2	6.5	.88	26.0	6.2	24.1	6.6	26.25	<.001	0.332
C5 (principled)	29.3	4.5	.52	30.2	4.0	28.2	4.8	50.54	<.001	0.461
Cooperativeness (C)	134.1	16.3	.85	137.6	15.5	129.4	16.3	68.82	<.001	0.538
ST1 (self-forgetfulness)	29.9	6.2	.70	29.7	6.4	30.2	6.1	1.50	.22	0.079
ST2 (transpersonal identification)	21.4	5.1	.70	21.1	5.0	21.7	5.2	1.76	.18	0.086
ST3 (spiritual acceptance)	19.6	6.5	.79	19.8	6.4	19.2	6.6	2.08	.14	0.094
Self-transcendence (ST)	70.8	14.4	.85	70.6	14.4	71.2	14.5	0.12	.73	0.022

Note. Gender differences are shown by multivariate analysis of covariance. TCI-R = Temperament and Character Inventory-Revised. <sup>a</sup> Cohen's *d* values were computed from the F-test output, which controls for the influence of age.

The correlations among the four dimensions of temperament and the three dimensions of character are summarized in Table 2. Only correlations greater than .20 were significant ( $p < .001$ ). Significant negative correlations associated NS with HA (-.29) and age (-.22) and HA with PS (-.29) and SD (-.45). Significant positive correlations associated RD with C (.51), PS with SD (.34), and SD with C (.41) and age (.27).

Separate factorial analyses were performed for temperaments and characters. Factors were retained as long as the  $i$ th eigenvalue from the actual data was greater than the  $i$ th eigenvalues from the random value. For the temperaments, principal component analysis identified four factors with eigenvalues greater than those computed by the parallel analysis (Table 3). Figure 1 displays the eigenvalues of the actual and simulated data. The eigenvalues for the first six factors were, respectively, 3.66, 2.66, 2.18, 1.13, 0.98, and 0.87. The four factors accounted for 23%, 17%, 14%, and 7% of the variance (60% cumulatively). The fourth factor is not readily differentiated from the fourth random factor in the parallel analysis (eigenvalue of 1.13 for real data vs. 1.11 for simulated data). The standardized factor loading following promax rotation showed that in the four-factor solution, PS, HA, RD, and NS factors were robust. The subscale NS1 loaded also negatively on Factor 2 (HA).

**TABLE 2:** Correlations Among Temperament and Character Scales of the TCI-R

	<i>NS</i>	<i>HA</i>	<i>RD</i>	<i>PS</i>	<i>SD</i>	<i>C</i>	<i>ST</i>
<b>NS</b>							
<b>HA</b>	<u>-.29</u>	—	—	—	—	—	—
<b>RD</b>	.13	-.01	—	—	—	—	—
<b>P</b>	-.15	<u>-.29</u>	.04	—	—	—	—
<b>SD</b>	-.15	<u>-.45</u>	.10	<u>.34</u>	—	—	—
<b>CO</b>	-.16	-.08	<u>.51</u>	.16	<u>.41</u>	—	—
<b>ST</b>	.07	-.09	.20	.12	-.12	.09	—
<b>Age</b>	<u>-.22</u>	-.10	-.12	.19	<u>.27</u>	.11	.09

Note. Underlined coefficients are significant at  $p < .001$ . TCI-R = Temperament and Character Inventory-Revised; NS = Novelty Seeking; HA = Harm Avoidance; RD = Reward Dependence; PS = Persistence; SD = Self-Directedness; C = Cooperativeness; ST = Self-Transcendence.

**TABLE 3:** Factor Structure of the Subscales of the Temperaments From the TCI-R After Promax Rotation

Temperaments	Factor 1	Factor 2	Factor 3	Factor 4
<b>NS1</b>	.20	-.46	.32	<u>.53</u>
<b>NS2</b>	-.17	-.08	-.07	<u>.74</u>
<b>NS3</b>	-.20	-.09	.29	<u>.60</u>
<b>NS4</b>	-.15	-.28	-.06	<u>.71</u>
<b>HA1</b>	-.16	<u>.80</u>	.09	-.16
<b>HA2</b>	-.16	<u>.80</u>	-.01	-.37
<b>HA3</b>	-.21	<u>.74</u>	-.29	-.17
<b>HA4</b>	-.39	<u>.68</u>	.03	-.12
<b>RD1</b>	.21	.27	<u>.56</u>	.06
<b>RD2</b>	.18	-.20	<u>.86</u>	.15
<b>RD3</b>	-.08	-.10	<u>.78</u>	.16
<b>RD4</b>	-.10	.17	<u>.52</u>	-.17
<b>PS1</b>	<u>.83</u>	-.21	.10	-.13
<b>PS2</b>	<u>.85</u>	-.31	.02	-.10
<b>PS3</b>	<u>.82</u>	-.17	.03	.10
<b>PS4</b>	<u>.86</u>	-.20	.04	-.23

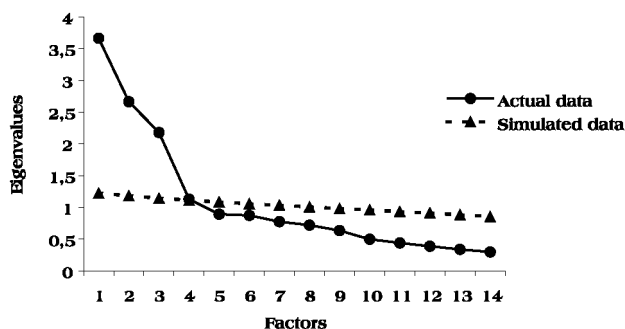
Note. Theoretically expected loadings are underlined. TCI-R = Temperament and Character Inventory-Revised; NS = Novelty Seeking; HA = Harm Avoidance; RD = Reward Dependence; PS = Persistence.

For the character scales, principal component analysis identified three factors with eigenvalues greater than those provided by the parallel analysis (Table 4). These accounted for 27%, 17%, and 13% of the variance (57% cumulatively).

Figure 2 shows the eigenvalues of the actual and simulated data. The eigenvalues for the first 6 factors were, respectively, 3.49, 2.24, 1.70, 0.89, 0.73, and 0.64. The standardized factor loading following promax rotation showed that C and ST factors were robust. For SD, SD1, SD2, SD3, and SD5, subscales loaded consistently, whereas SD4 subscale had only very weak loading (.29). In fact, the SD4 subscale loaded more strongly on Factor 1 (C).

Table 5 presents the test-retest reliabilities of TCI-R scales and subscales. All ICCs were statistically significant. At 2 weeks, all the scales and subscales showed good to very good reliabilities, with ICCs ranging from .69 to .94. For the scales, ICCs were very high (.81 for ST to .94 for PS). For the subscales, only two subscales had ICC values lower than .77 (C2 and C5). The one-way repeated measures ANOVA showed that HA and ST were lower at the retest, whereas NS was higher. Concerning the subscales, HA1, HA2, HA3, RD1, and all the ST subscales were lower at retest, although NS2, NS4, and SD5 were higher. In contrast, reliability of TCI-R scales and subscales at 6 months was a little weaker, particularly for the NS scale. ICCs ranged from .46 to .84. For the scales, the lowest fidelity was observed for NS (.53) and the highest for SD (.84). For the subscales, lowest reliabilities were observed for NS4 (.58), HA2 (.59), HA4 (.46), RD4 (.47), PS2 (.56), SD1 (.55), and C5 (.59). The oneway repeated measures ANOVA showed that only NS3 and HA4 were higher at retest, whereas NS1, RD1, and RD4 were lower.

**FIGURE 1:** Scree plot of eigenvalues derived from the principal component analysis (promax rotation) computed on the temperament's subscales of this data (actual data). Eigenvalues of the simulated data are obtained from parallel analysis. The inspection of the two curves justify the extraction of four factors: The first 4 factors derived from the actual data exhibit higher eigenvalues than the first 4 factors obtained by parallel analysis, but the fifth factor from the actual data has lower eigenvalue than the fifth factor given by parallel analysis.

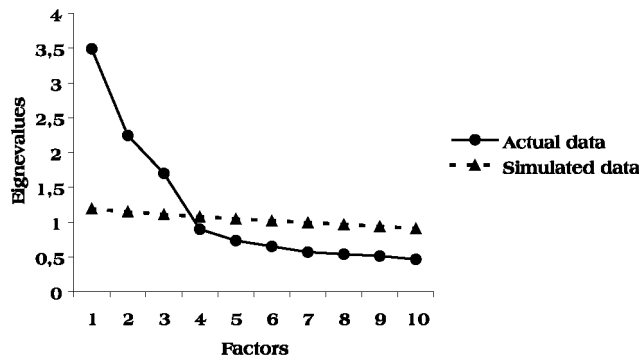


**TABLE 4:** Factor Structure of the Character Subscales From the TCI-R After Promax Rotation

Characters	Factor 1	Factor 2	Factor 3
SD1	.33	<u>.74</u>	-.20
SD2	.27	<u>.71</u>	.18
SD3	.18	<u>.80</u>	-.06
SD4	.56	<u>.29</u>	-.34
SD5	.23	<u>.80</u>	-.01
C1	<u>.74</u>	.31	.14
C2	<u>.61</u>	.20	.41
C3	<u>.78</u>	.32	.07
C4	<u>.75</u>	.18	-.03
C5	<u>.61</u>	.10	-.03
ST1	-.11	-.06	<u>.80</u>
ST2	.16	.09	<u>.80</u>
ST3	.07	-.08	<u>.75</u>

Note. Theoretically expected loadings are underlined. TCI-R = Temperament and Character Inventory-Revised; SD = Self-Directedness; C = Cooperativeness; ST = Self-Transcendence.

**FIGURE 2:** Scree plot of eigenvalues derived from the principal component analysis (promax rotation) computed on the character's subscales of this data (actual data). Eigenvalues of the simulated data are obtained from parallel analysis. The inspection of the two curves justify the extraction of three factors: the first 3 factors derived from the actual data exhibit higher eigenvalues than the first 3 factors obtained by parallel analysis, but the fourth factor from the actual data has lower eigenvalue than the fourth factor given by parallel analysis.



## DISCUSSION

Since the introduction of the TCI-R by Cloninger in 1999, few data have been available in the literature (Brändström et al., 2003; Gutierrez-Zotes et al., 2004; Pélissolo et al., in press). In this study, we provided further psychometric properties of the French version of the TCI-R administered to a Belgian sample. Overall, the factorial structure of the TCI-R was good. For the temperaments, PS is now a very robust factor, with factor loadings higher than 0.80 for the four subscales. This point is of interest because PS contains 35 items in the revised version of the TCI, whereas it had only 8 items in the TCI. The inclusion of PS subscales is therefore a useful extension. This is confirmed by the excellent alpha coefficient (0.90) and excellent test-retest reliability (0.94 at 2 weeks). The factor of HA was also robust, and its internal consistency is similar to PS value (0.90). The RD factor was less strong, mainly because RD1 and RD4 displayed poorer factor loadings. In contrast, the internal consistency of this scale was good (0.81). Again, NS was less robust; NS1 and NS3 exhibited lower factor loadings, and in consequence, NS showed the lowest internal consistency as revealed by its lower alpha coefficient (0.78). It should be noted that this factor is smaller and less well defined than the others and that it is not readily differentiated from the random factor in parallel analysis. However, its composition is relatively consistent with theoretical expectations, which argues for its retention. The factorial structure found here for the character dimensions is confirmed, except for the SD subscale SD4 that appeared to be less specific. The inspection of Table 4 shows that SD4 loaded positively (.56) on the C scale. This means that self-acceptance (SD4) has an important relationship with the C dimension. The theoretical issue about the relationship of SD4 to C and SD is discussed in depth elsewhere (Cloninger, 2004). Briefly, character development involves nonlinear adaptive changes among all aspects of personality and cannot be adequately represented in terms of linear dimensions alone (Cloninger, 2004). Self-acceptance (SD4) measures the emotional aspects of SD that are associated with C, but SD4 develops more slowly than C. In a relatively young sample such as that studied here, SD4 was not well developed.

Brändström et al. (2003) did not describe the factorial analysis results of the TCI-R, but they provided comparisons of the factor structures of the TCI and the TCI-R based on principal axis factor analysis with oblimin rotations that have shown congruence coefficients mostly above .90, suggesting high similarity in the factor structures of both versions of the questionnaire. Brändström et al. (2003) reported that Cronbach alpha coefficients for the revised version were on average higher than those found for the TCI in both Swedish and German versions (i.e., 0.60 to 0.87 for the TCI and 0.82 to 0.90 for the TCI-R) as well as higher than those found in this study. Pélissolo et al. (in press) found that the factorial structure of the TCI-R was comparable to the one reported in this study. For the temperaments, PS was clearly identified as well as RD, and NS1 loaded negatively on the HA factor. For the characters, Pélissolo et al. demonstrated also that SD4 loaded positively on the C factor. Similarly to our results, Pélissolo et al. found high Cronbach alpha coefficients for all the dimensions (from 0.80 to 0.92). High Cronbach alpha coefficients were also reported by Gutierrez-Zotes et al. (2004; from 0.77 for NS to 0.85 to SD), but these authors did not perform factorial analysis.

Interestingly, the factor analysis results of this study are consistent with those found with the TCI by Pélissolo and Lépine (1997) and almost identical to those reported by Cloninger et al. (1993) and Hansenne et al. (2001). Indeed, previous TCI studies reported that subscale NS 1 (exploratory excitability) loaded more on the HA factor

(Hansenne et al., 2001; Lépine, Pélissolo, Téodorescu, & Téhérani, 1994; Pélissolo & Lépine, 1997). In this study, we also found this association. Moreover, within the TCI, SD4 had already shown poor factor loading on SD (Hansenne et al., 2001).

**TABLE 5:** Test-Retest Reliabilities at 2 Weeks<sup>a</sup> and at 6 Months<sup>b</sup>

TCI-R Dimensions	2 Weeks					6 Months						
	Test	Retest	ICC	F(1, 77)	P	d <sup>c</sup>	Test	Retest	ICC	F(1, 39)	P	d <sup>c</sup>
NS1 (exploratory excitability)	32.9	32.8	.82	0.17	.67	0,047	30.9	29.4	.65	6.01	.02	0.393
NS2 (impulsiveness)	24.2	25.0	.83	6.97	<.01	0,301	23.4	24.1	.60	0.87	.35	0.149
NS3 (extravagance)	26.6	26.9	.89	0.19	.65	0,050	24.1	26.3	.66	9.61	<.01	0.496
NS4 (disorderliness)	19.3	20.1	.81	9.76	<.01	0,356	18.9	19.1	.58	0.09	.75	0.048
Novelty seeking (NS)	103.2	104.8	.92	8.01	<.01	0,323	97.2	96.7	.53	0.47	.82	0.110
HA1 (anticipatory worry)	32.8	31.6	.85	9.73	<.01	0,355	32.7	32.9	.68	0.07	.79	0.042
HA2 (fear of uncertain)	22.8	22.2	.85	5.14	.02	0,258	24.4	24.4	.59	0.01	.91	0.016
HA3 (shyness with strangers)	22.3	21.6	.94	7.16	<.01	0,305	23.4	24.4	.69	2.27	.13	0.241
HA4 (fatigability)	21.3	21.4	.81	0.09	.75	0,034	21.0	23.0	.46	7.01	.01	0.424
Harm avoidance (HA)	99.3	96.7	.91	12.02	<.01	0,395	101.4	104.7	.72	2.66	.11	0.261
RD1 (sentimentality)	30.2	29.5	.78	6.52	.01	0,291	30.5	29.0	.64	7.40	<.01	0.436
RD2 (open to warm communication)	35.7	35.6	.90	0.01	.93	0,011	34.4	34.0	.66	0.23	.63	0.077
RD3 (attachment)	21.1	21.2	.86	0.01	.93	0,011	20.1	20.1	.74	0.01	.96	0.016
RD4 (dependence)	20.5	20.0	.79	3.32	.07	0,208	17.1	19.8	.47	24.3	<.001	0.789
Reward dependence (RD)	107.4	106.4	.89	2.99	.08	0,197	102.3	103.0	.77	0.20	.65	0.072
PS1 (eagerness of effort)	29.4	29.6	.81	0.08	.77	0,032	29.1	29.5	.78	0.59	.44	0.123
PS2 (work hardened)	27.2	27.1	.85	0.18	.66	0,048	26.7	26.9	.56	0.07	.78	0.042
PS3 (ambitious)	33.2	33.5	.89	0.85	.36	0,105	33.2	33.5	.73	0.14	.70	0.060
PS4 (perfectionist)	27.2	27.3	.89	0.20	.64	0,051	26.3	26.1	.64	0.04	.83	0.032
Persistence (PS)	117.0	117.6	.94	0.26	.60	0,058	115.4	116.1	.77	0.12	.72	0.055
SD1 (responsibility)	29.8	30.2	.79	0.69	.40	0,095	28.5	29.2	.55	0.96	.33	0.157
SD2 (purposefulness)	22.5	22.5	.82	0.01	.92	0,011	21.4	21.5	.81	0.05	.82	0.036
SD3 (resourcefulness)	17.9	17.5	.86	2.69	.10	0,187	17.3	17.3	.72	0.02	.89	0.023
SD4 (self-acceptance)	30.9	31.3	.89	0.70	.41	0,095	30.8	29.8	.63	1.14	.29	0.171
SD5 (congruent second nature)	36.8	37.6	.87	4.93	.03	0,253	37.3	36.5	.62	1.01	.31	0.161
Self-directedness (SD)	137.9	139.2	.93	2.08	.15	0,164	135.4	134.4	.76	0.28	.59	0.085
C1 (social acceptance)	30.7	30.7	.77	0.11	.74	0,038	30.0	30.3	.73	0.26	.61	0.082
C2 (empathy)	19.3	19.3	.76	0.02	.89	0,016	18.5	18.8	.65	0.70	.40	0.134
C3 (helpfulness)	30.8	30.7	.83	0.18	.66	0,048	30.3	29.9	.74	0.61	.43	0.125
C4 (compassion)	24.8	25.0	.92	0.33	.56	0,065	24.7	23.9	.82	1.10	.29	0.168
C5 (principled)	29.9	29.2	.69	3.77	.06	0,221	29.4	28.7	.59	1.57	.21	0.201
Cooperativeness (C)	135.5	134.9	.91	0.38	.53	0,070	132.9	131.7	.84	0.51	.47	0.114
ST1 (self-forgetfulness)	30.2	28.7	.77	22.94	<.01	0,546	30.0	30.1	.71	0.02	.87	0.023
ST2 (transpersonal identification)	20.9	19.9	.80	11.85	<.01	0,392	19.9	20.2	.71	0.22	.63	0.075
ST3 (spiritual acceptance)	20.0	18.9	.85	14.13	<.01	0,428	20.2	20.6	.63	0.21	.64	0.073
Self-transcendence (ST)	71.1	67.5	.81	29.9	<.01	0,623	70.1	70.9	.77	0.24	.62	0.078

Note. TCI-R = Temperament and Character Inventory-Revised; ICC = intraclass correlation coefficients;  $d$  = Cohen's  $d$  effect size. <sup>a</sup> $N = 78$ . <sup>b</sup> $N = 40$ . <sup>c</sup>Cohen's  $d$  values were computed from the  $F$  test output, which takes into account the correlation between baseline and retest scores in this within-subjects design.

Test-retest reliability of TCI-R scales and subscales at 2 weeks was very satisfactory, but the ICCs obtained at 6 months were merely fair to good, a result suggesting that consistency of the scales over time is perhaps a weakness of the TCI-R and more particularly for the NS scale. This was confirmed by higher differences in effect size for test-retest at 6 months. Effects sizes at 2 weeks were only moderate (0.20) for ST and ST subscales ST1 and ST2. Nevertheless, the limited sample size for the retest at 6 months could have contributed to the weak ICC coefficients.



In contrast to our results, Pélissolo et al. (in press) provided very good test-retest data on 21 psychiatric patients (3-week interval) and 15 healthy participants (1-week interval) with the TCI-R. The ICC coefficients ranged from .82 to .93, and the ICC coefficients for NS were .91 and .93, respectively, in the clinical sample and in healthy participants. Unfortunately, neither Brändström et al. (2003) nor Gutierrez-Zotes et al. (2004) have computed test-retest coefficients. However, Cloninger et al. (1991) reported 6-month test-retest correlations from a previous instrument (Tridimensional Personality Questionnaire) that were comparable to those found in this study (i.e., .79 for HA, .70 for RD, and .76 for NS). The main difference appears for the NS dimension. Pélissolo and Lépine (1997) reported very high ICC coefficients for the TCI scales among a sample of 65 undergraduate students for an interval of 32 days between test and retest. They provided only the scales' ICCs, and no information was available concerning subscales. In general, 2-week ICC values of this study were similar to those reported by Pélissolo and Lépine (1997).

The results show that gender influences significantly some dimensions of the TCI-R. Specifically, women exhibited higher scores on HA, RD, and C than men. Pélissolo et al. (in press) found higher TCI-R scores on RD, SD, and C among women. Again, with the TCI-R, Gutierrez-Zotes et al. (2004) reported that women scored higher on HA, RD, and C, although men scored higher on PS. Approximately the same effects were observed for the TCI among 322 representative Belgian participants (Hansenne et al., 2001), with the exception that in the TCI study, SD was lower in women. Moreover, Cloninger et al. (1993) showed that women had higher scores on the C scale and also the ST subscale ST3. With the French version of the TCI, Pélissolo and Lépine (1997) also found that women tended to have higher scores on HA and RD. One may question these gender differences. Concerning higher HA scores in women found in this study and in the study of Gutierrez-Zotes et al. (2004), it is usually reported that women are on average more anxious than men, and several epidemiological studies have reported a greater prevalence of anxiety disorders among women (Fullerton et al., 2001; Sheikh, Leskin, & Klein, 2002). Moreover, it was suggested from a twin study that the increased prevalence of anxiety disorders in women may occur because anxiety sensitivity could be heritable in women (Jang, Stein, Taylor, & Livesley, 1999). An alternative explanation may be that men underreport depressive and anxiety symptoms because it is socially undesirable for them to admit mental health problems (Vredenburg, Kramers, & Flett, 1986).

This study showed that the TCI-R scales were weakly related among themselves. Pélissolo et al. (in press) found also weak correlations among scales of the TCI-R. Moreover, the relationships are consistent with those reported by Cloninger et al. (1993) and by Hansenne et al. (2001) for the TCI. The highest correlation in this study was observed between RD and C, suggesting that individuals characterized by attachment, dependency, and sentimentality exhibit good empathy, compassion, and social acceptance. This higher association was probably caused by the addition of a RD2 subscale (openness to warm communication) within the RD scale. The second highest correlation found here related SD with HA. In the study conducted by Pélissolo et al. (YEAR), SD was highly negatively correlated with HA (-.60). This relation can mean that the anxious participants had more difficulties choosing goals and personal values and that they did not accept themselves. In addition, several studies have demonstrated that some dimensions, especially the NS dimension, varied with age (Cloninger et al., 1991; Hansenne et al., 2001; Le Bon, Staner, Tecco, Pull, & Pelc, 1998; Lépine et al., 1994). Moreover, SD and C scores were reported to have a strong correlation with age (Cloninger et al., 1993). This study confirms the relationship between both NS and SD scales and age. Last, SD and C were associated. This point confirms that these two scales are interrelated developmentally as a person matures and validates the effectiveness of adding them to create a maturity score used to determine the presence or absence of personality disorders (Bayon et al., 1996; Cloninger & Svrakic, 1997).

Finally, it is interesting to observe that the mean scores of the main dimensions obtained in our sample were consistent with those reported by the other European studies that used the TCI-R (Brändström et al., 2003; Gutierrez-Zotes et al., 2004; Pélissolo et al., in press). This means that German, Spanish, Sweden, French, and Belgian participants have approximately the same mean scores on the dimensions proposed by Cloninger and colleagues (Cloninger et al., 1993) and that cultural differences found in these countries do not influence mean personality dimensions among the population.

Several limitations of this study should be acknowledged. First, although the sample was beyond the typical college student sample, it comprised more women (546) than men (412). The sex disparity within the sample between students and relatives was even more problematic. Indeed, there were more men in the relative group (374) than in the student group (68), whereas for women, the two groups were comparable (282 vs. 264). This occurred because there were more female students in psychology courses in Belgium, and so to compensate this bias, we recruited more men in the relative group. However, we have no information about the proportion of students within the relative group. Second, and more important, this validation study was limited to the TCI-R data, and we have no external criteria or other personality dimensions provided by other tests to relate with the

results. In short, the study lacks evidence of convergent or discriminant validity. Further studies should be conducted in which the TCI-R is completed in conjunction with other measures of personality. More particularly, the relationships between the TCI-R and the Five-factor model inventory (Revised NEO Personality Inventory [NEO-PI-R]) would be of interest because some studies have already assessed the relationships between the previous version of the questionnaire (TCI) and the NEO-PI-R (De Fruyt, Van De Wiele, & Van Heeringen, 2000; MacDonald & Holland, 2002).

In conclusion, this study confirms that the revised version of the TCI displays relatively good psychometric properties. Particularly, the PS dimension appears to be a robust dimension that is well measured by the TCI-R. The self-acceptance subscale SD4 was weakly correlated with the other SD subscales in this young sample, so its development should be further examined in a longitudinal study or in an older population. Globally, the results presented here give further support of the psychobiological model proposed by Cloninger and his colleagues (Cloninger et al., 1993). However, because temperament scales and character scales are conceptualized as having nonlinear rather than linear links, it is difficult to incorporate them in a single integrated factor model, and consequently, this point constitutes a psychometric limitation of the TCI-R.

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