# A Bayesian approach for parameter identification in elastoplasticity

#### Hussein Rappel, Lars Beex, Jack Hale, Stéphane Bordas

hussein.rappel@uni.lu

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### Introduction



### Introduction



#### Error minimisation

# Least squares method (conventional approach): $\sigma = E\epsilon$ $J = \frac{1}{2} \sum_{i=1}^{N} (\sigma_i - E\epsilon_i)^2$ $\overline{E} = \underset{E}{\operatorname{argmin}} J(E)$

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#### Frequentist inference



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### Frequentist inference





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#### Frequentist inference





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#### Method of maximum likelihood (ML):

$$\pi(\sigma^{m}|E) = \frac{1}{\sqrt{2\pi}S_{noise}} \exp\left(-\frac{(\sigma^{m}-E\epsilon)^{2}}{2S_{noise}^{2}}\right)$$

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Method of maximum likelihood (ML):

$$\pi(\sigma^{m}|E) = \frac{1}{\sqrt{2\pi}S_{noise}} \exp\left(-\frac{(\sigma^{m}-E\epsilon)^{2}}{2S_{noise}^{2}}\right)$$

and for M measurments:

$$\boldsymbol{\sigma}^{m} = \begin{bmatrix} \sigma_{1}, \sigma_{2}, \cdots, \sigma_{M} \end{bmatrix}$$

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### Bayesian inference



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### **Bayesian** inference



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$$\sigma^{m} = E\epsilon + \Omega$$

$$\Omega \sim \pi_{noise}(\omega)$$

$$\sigma^{m} = 0$$

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#### **Bayes' formula:**

$$\pi(E|\sigma^m) = \frac{\pi(E)\pi(\sigma^m|E)}{\pi(\sigma^m)}$$

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#### Bayes' formula:

$$\pi(E|\sigma^m) = \frac{\pi(E)\pi(\sigma^m|E)}{\pi(\sigma^m)} \implies \frac{\pi(E)\pi(\sigma^m|E)}{C}$$

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**Bayes' formula:** 

$$\pi(E|\sigma^m) = \frac{\pi(E)\pi(\sigma^m|E)}{\pi(\sigma^m)} \implies \frac{\pi(E)\pi(\sigma^m|E)}{C}$$

$$\pi(E|\sigma^m)\propto \pi(E)\pi(\sigma^m|E)$$

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$$\pi(E|\sigma^m) \propto \exp\left(-\frac{(E-\overline{E})^2}{2S_E^2}\right) \exp\left(-\frac{(\sigma^m - E\epsilon)^2}{2S_{noise}^2}\right)$$

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$$\pi(E|\sigma^m) \propto \exp\Big(-rac{(E-\overline{E})^2}{2S_E^2}\Big) \exp\Big(-rac{(\sigma^m-E\epsilon)^2}{2S_{noise}^2}\Big)$$

#### and for M measurements:

$$\begin{aligned} \pi(E|\boldsymbol{\sigma}^{m}) \propto \prod_{i=1}^{M} \exp\left(-\frac{(E-\overline{E})^{2}}{2S_{E}^{2}}\right) \exp\left(-\frac{(\sigma_{i}^{m}-E\epsilon_{i})^{2}}{2S_{noise}^{2}}\right) \\ = \left[\sigma_{1}, \sigma_{2}, \cdots, \sigma_{M}\right] \end{aligned}$$

 $\sigma^m$ 

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$$\pi(E|\boldsymbol{\sigma}^m)\propto \exp\Big(-rac{(E-\mu)^2}{2S_{post}^2}\Big)$$

with 
$$\mu = f(\sigma_i, \overline{E}, S_{nosie}, S_E, \epsilon_i)$$
  
 $S_{post} = f(\sigma_i, \overline{E}, S_{nosie}, S_E, \epsilon_i)$ 

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## Bayesian inference: linear elastic-perfectly plastic model identification



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### Bayesian inference: linear elastic-linear hardening model identification





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### Bayesian inference: linear elastic-nonlinear hardening model identification



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# Bayesian inference: Young's modulus identification with double uncertainty



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# Bayesian inference: Young's modulus identification with double uncertainty



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# Bayesian inference: linear elastic-perfectly plastic with double uncertainty



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# Bayesian inference: linear elastic-linear hardening with double uncertainty



# Bayesian inference: linear elastic-nonlinear hardening with double uncertainty



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- 'Closed form expression' of the posterior for:
  - linear elasticity,
  - elastoplasticity with perfect plasticity,
  - elastoplasticity with linear hardening, and
  - elastoplasticity with nonlinear hardening.

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- 'Closed form expression' of the posterior for:
  - linear elasticity,
  - elastoplasticity with perfect plasticity,
  - elastoplasticity with linear hardening, and
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- The results of BI cannot directly be compared to those of the least squares method.

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- 'Closed form expression' of the posterior for:
  - linear elasticity,
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- 'Closed form expression' of the posterior for:
  - linear elasticity,
  - elastoplasticity with perfect plasticity,
  - elastoplasticity with linear hardening, and
  - elastoplasticity with nonlinear hardening.
- The results of BI cannot directly be compared to those of the least squares method.
- The selected prior distribution has a significant effect on the results.
- BI leads to a distribution for the considered parameters, however the resulting distribution do *not* reflect the heterogeneity of the material parameters.

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#### Future work



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### The End

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